



# Nonparametric estimation of the local Hurst function of multifractional Gaussian processes

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# Nonparametric estimation of the local Hurst function of multifractional Gaussian processes

Joint paper with Donatas Surgailis (Lithuania)

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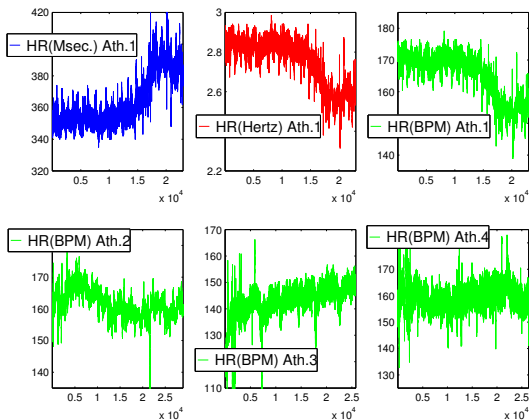
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- 1 Introduction
- 2 A new nonparametric estimator of the local Hurst function
  - First definition, assumptions and limit theorems
  - Second definition and limit theorems
  - Case of the General multifractional Brownian motion
- 3 Numerical comparison with the quadratic variations estimators
  - Definition and limit theorems
  - Numerical comparisons

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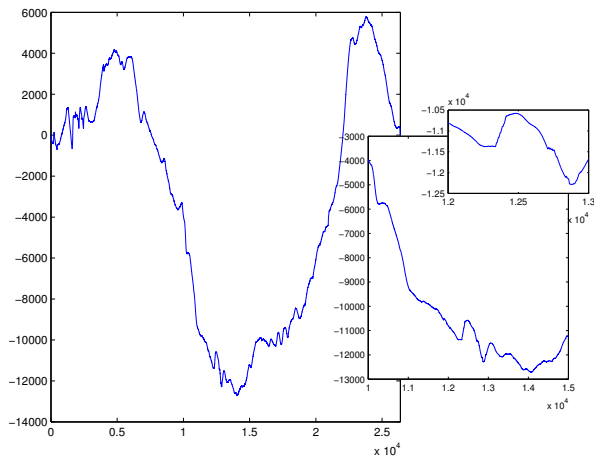
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# An athletic example...



Heart beat of athletes during Marathon de Paris 2004

## A fact...



Successive zooms on the aggregated series of one athlete

# The fractional Brownian motion

- $B^H = \{B_t^H, t \in \mathbb{R}\}$  **fractional Brownian motion** (FBM) with  $H \in [0, 1]$ :

$B^H$  is a Gaussian centered process with stationary increments and  $\text{Var}(B_t^H) = \sigma^2 |t|^{2H}$ .

## Property

*A Gaussian process  $X$  having stationary increments  **$H$ -selfsimilar***

$$\iff X \text{ F.B.M. with parameter } H$$

- A trajectory of  $B^H$  is a.s.  **$\alpha$ -Hölderian** for any  $\alpha < H$ :  
 $\implies H$  measures the **local smoothness** of  $B^H$  also called **Hurst parameter**

# Two trajectories of FBM

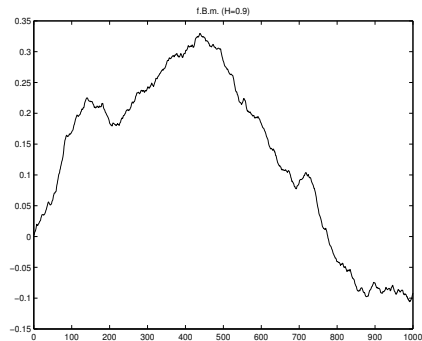
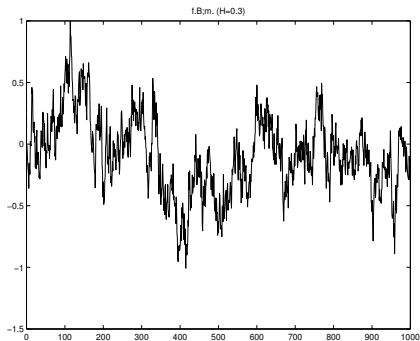


Figure: Trajectories of FBM with  $H = 0.3$  (left) and  $H = 0.9$  (right)



# Other definitions of FBM

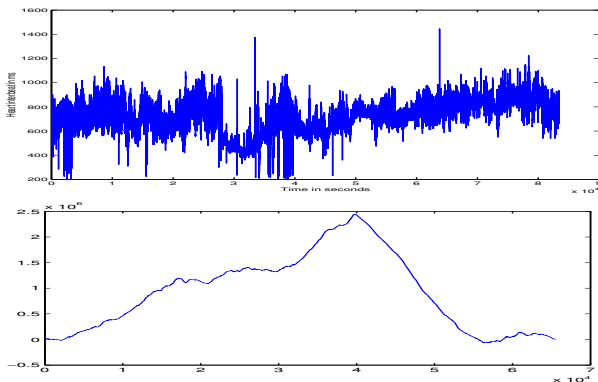
- Harmonizable representation:

$$B_t^H = C_1(H) \sigma^2 \int_{\mathbb{R}} \frac{e^{it\xi} - 1}{|\xi|^{2H+1}} \widehat{W}(d\xi) \quad t \in \mathbb{R}$$

- Temporal representation:

$$B_t^H = C_2(H) \sigma^2 \int_{\mathbb{R}} ((t-u)_+^{H-1/2} - (-u)_+^{H-1/2}) W(du) \quad t \in \mathbb{R}$$

# Another example



Heartbeats during 24h.

⇒  $H$  depending on  $t$ !

# Multifractional Brownian motion

Two first versions:

- **Harmonizable representation** (Benassi *et al.*, 1997):

$$X_t = C_1(H(t)) \sigma^2 \int_{\mathbb{R}} \frac{e^{it\xi} - 1}{|\xi|^{2H(t)+1}} \widehat{W}(d\xi) \quad t \in \mathbb{R}$$

- **Temporal representation** (Peltier et Lévy-Véhel, 1995):

$$X_t = C_2(H(t)) \sigma^2 \int_{\mathbb{R}} ((t-u)_+^{H(t)-1/2} - (-u)_+^{H(t)-1/2}) W(du) \quad t \in \mathbb{R}$$

# Aims

From an observed trajectory  $(X_{\frac{1}{n}}, X_{\frac{2}{n}}, \dots, X_1)$ ,

- Define a new non-parametric estimator  $H(\cdot)$ ;
- Asymptotic properties of this estimator.
- Compare this estimator with the well-known quadratic variations estimator.

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# First definition

For  $t_0 \in (0, 1)$  and  $\alpha \in (0, 1)$ , define:

$$\widehat{H}_{n,\alpha}^{(IR)}(t_0) := \Lambda^{-1} \left( \frac{1}{2n^{1-\alpha}} \sum_{k=[nt_0-n^{1-\alpha}]}^{[nt_0+n^{1-\alpha}]} \frac{|\Delta_n^k X + \Delta_n^{k+1} X|}{|\Delta_n^k X| + |\Delta_n^{k+1} X|} \right)$$

with

- $\Delta_n^k X = X_{\frac{k+2}{n}} - 2X_{\frac{k+1}{n}} + X_{\frac{k}{n}}$
- $\Lambda(h) = \mathbb{E} \left[ \frac{|\Delta_1^0 B^h + \Delta_1^1 B^h|}{|\Delta_1^0 B^h| + |\Delta_1^1 B^h|} \right]$  for  $h \in (0, 1)$   
 $( = \frac{1}{\pi} \arccos(-\rho_2(h)) + \frac{1}{\pi} \sqrt{\frac{1+\rho_2(h)}{1-\rho_2(h)}} \log \left( \frac{2}{1+\rho_2(h)} \right) \text{ with } \rho_2(h) = \frac{-3^{2h}+2^{2h+2}-7}{8-2^{2h+1}} )$ .

$\Rightarrow$  Explanation...

# Multifractional Gaussian processes?

**Assumptions:**  $X = (X_t)_t$  is a centered Gaussian process such as:

- (A) $_{\kappa}$**  There exist  $\eta$ -Hölderian functions  $0 < H(t) < 1$  and  $c(t) > 0$  for  $t \in (0, 1)$  such that for any  $0 < \varepsilon < 1/2$  and  $j \in \mathbb{Z}$ ,

$$\max_{[n\varepsilon] \leq k \leq [(1-\varepsilon)n]} n^{\kappa} \left| \frac{\text{Cov}(\Delta_n^k X, \Delta_n^{k+j} X)}{\text{Cov}(\Delta_n^k B^{H(k/n)}, \Delta_n^{k+j} B^{H(k/n)})} - c\left(\frac{k}{n}\right) \right| \xrightarrow{n \rightarrow \infty} 0.$$

- (B) $_{\alpha}$**  There exist  $C > 0$ ,  $\gamma > 1/2$  and  $0 \leq \theta < \gamma/2$  such that for any  $n \in \mathbb{N}^*$  and  $1 \leq k, k' < n - q$

$$\left| \text{Cor}(\Delta_n^k X, \Delta_n^{k'} X) \right| \leq C n^{(1-\alpha)\theta} (|k' - k| \wedge n^{1-\alpha})^{-\gamma}.$$

# Limit theorems for multifractional Gaussian processes

## Theorem

Under Assumptions **(A) $_{\kappa}$**  and **(B) $_{\alpha}$** , for all  $t_0 \in (0, 1)$ ,

- If  $0 < \alpha < \frac{\gamma-2\theta}{2(\gamma-\theta)}$ ,

$$\widehat{H}_{n,\alpha}^{(IR)}(t_0) \xrightarrow[n \rightarrow \infty]{a.s.} H(t_0).$$

- If  $\kappa \geq \mu$  and  $\frac{2\gamma-2\theta}{3\gamma-2\theta+4\gamma(\eta \wedge 2)} \leq \alpha < \frac{2\gamma-2\theta}{3\gamma-2\theta}$  then for any  $\epsilon > 0$

$$\sup_{\epsilon < t < 1-\epsilon} |\widehat{H}_{n,\alpha}^{(IR)}(t) - H(t)| = O_p(n^{-\mu}).$$

- If  $\kappa \geq \mu_1$  and  $\frac{\gamma-2\theta}{3\gamma-2\theta+4\gamma(\eta \wedge 2)} \leq \alpha < \frac{\gamma-2\theta}{3\gamma-2\theta}$  then for any  $\epsilon > 0, \delta > 0$

$$\sup_{\epsilon < t < 1-\epsilon} |\widehat{H}_{n,\alpha}^{(IR)}(t) - H(t)| = O(n^{-(\mu_1-\delta)}) \quad a.s.$$



# Central Limit theorem for multifractional Gaussian processes

## Theorem

Let  $Z = (Z(t))_{t \in (0,1)}$  be a zero-mean Gaussian process satisfying  $(\mathbf{A})_\kappa$  and  $(\mathbf{B})_\alpha$ , with  $\alpha > \frac{1}{1+2(\eta \wedge 2)}$ ,  $\kappa \geq \frac{1-\alpha}{2}$  and  $\theta = 0$ . Then for  $0 < t_1 < \dots < t_u < 1$ ,

$$\sqrt{2n^{1-\alpha}} \left( \widehat{H}_{n,\alpha}^{(IR)}(t_i) - H(t_i) \right)_{1 \leq i \leq u} \xrightarrow[n \rightarrow \infty]{\mathcal{D}} \left( W^{(IR)}(t_i) \right)_{1 \leq i \leq u},$$

where  $W^{(IR)}(t_i)$ ,  $i = 1, \dots, u$  are independent centered Gaussian r.v.'s such as

$$\mathbb{E}[W^{(IR)}(t_i)]^2 := \left[ \frac{\partial}{\partial x} (\Lambda_2)^{-1} (\Lambda_2(H(t_i))) \right]^2 \sigma^2(H(t_i))$$

where 
$$\sigma^2(H) := \sum_{k \in \mathbb{Z}} \text{Cov} \left( \psi(\Delta_1^0 B_H, \Delta_1^1 B_H), \psi(\Delta_1^k B_H, \Delta_1^{k+1} B_H) \right).$$

# Proof: based on a CLT of triangular arrays of functional of Gaussian vectors

## Theorem (Bardet et Surgailis, 2012)

Let  $(\mathbf{Y}_n(k))_{1 \leq k \leq n, n \in \mathbb{N}}$  be a triangular array of standard Gaussian  $\mathbb{R}^\nu$ -vectors.

- For  $m \geq 1$ , there exists  $\rho : \mathbb{N} \rightarrow \mathbb{R}$  such as for  $1 \leq p, q \leq \nu$ ,

$$\forall(j, k), \quad \left| \mathbb{E} Y_n^{(p)}(j) Y_n^{(q)}(k) \right| \leq |\rho(j - k)| \quad \text{with} \quad \sum_{j \in \mathbb{Z}} |\rho(j)|^m < \infty;$$

- For  $\tau \in [0, 1]$  and  $J \in \mathbb{N}^*$ , with  $(\mathbf{W}_\tau(j))_{j \in \mathbb{Z}}$  a stationary Gaussian process

$$(\mathbf{Y}_n([n\tau] + j))_{-J \leq j \leq J} \xrightarrow[n \rightarrow \infty]{\mathcal{D}} (\mathbf{W}_\tau(j))_{-J \leq j \leq J};$$

- If  $\tilde{f}_{k,n} \in \mathbb{L}_0^2(\mathbf{X})$  ( $n \geq 1, 1 \leq k \leq n$ ) with Hermite rank  $\geq m$  and  $\exists \tilde{\phi}_\tau, \tau \in [0, 1]$

$$\text{such as} \quad \sup_{\tau \in [0, 1]} \|\tilde{f}_{[\tau n], n} - \tilde{\phi}_\tau\|^2 = \sup_{\tau \in [0, 1]} \mathbb{E}(\tilde{f}_{[\tau n], n}(\mathbf{X}) - \tilde{\phi}_\tau(\mathbf{X}))^2 \xrightarrow[n \rightarrow \infty]{} 0.$$

## Theorem

Then with  $\sigma^2 = \int_0^1 d\tau \left( \sum_{j \in \mathbb{Z}} \mathbb{E}[\tilde{\phi}_\tau(\mathbf{W}_\tau(0)) \tilde{\phi}_\tau(\mathbf{W}_\tau(j))] \right) < \infty$ ,

$$n^{-1/2} \sum_{k=1}^n \tilde{f}_{k,n}(\mathbf{Y}_n(k)) \xrightarrow[n \rightarrow \infty]{\mathcal{D}} \mathcal{N}(0, \sigma^2).$$

# Second definition

For  $t, \alpha \in (0, 1)$  and  $j = 1, \dots, p$ ,

$$\hat{H}_{n,\alpha,j}^{(IR)}(t_0) := \Lambda_j^{-1} \left( \frac{1}{2n^{1-\alpha}} \sum_{k=[nt_0-n^{1-\alpha}]}^{[nt_0+n^{1-\alpha}]} \frac{|\Delta_n^k X + \Delta_n^{k+j} X|}{|\Delta_n^k X| + |\Delta_n^{k+j} X|}, \right).$$

## Theorem

Under Assumptions  $(\mathbf{A})_\kappa$  and  $(\mathbf{B})_\alpha$ , for all  $t_0 \in (0, 1)$ ,

$$n^{(1-\alpha)/2} (\hat{H}_{n,\alpha,j}^{(IR)}(t_0) - H(t_0))_{1 \leq j \leq p} \xrightarrow[n \rightarrow \infty]{\mathcal{D}} \mathcal{N}(0, \Gamma(H(t_0)))$$

## Second definition (end)

Let  $\widehat{\Gamma} := \Gamma(\widehat{H}_{n,\alpha}^{(IR)})$ .

A new nonparametric estimator of  $H(\cdot)$  using **Pseudo-Generalized least squares**:

$$\widetilde{H}_{n,\alpha}^{(IR)}(t_0) := (\mathbb{1}_p' \widehat{\Gamma}^{-1} \mathbb{1}_p)^{-1} \mathbb{1}_p' \widehat{\Gamma}^{-1} (\widehat{H}_{n,\alpha,j}^{(IR)}(t_0))_{1 \leq j \leq p}.$$

### Theorem

Under Assumptions **(A) $_{\kappa}$**  and **(B) $_{\alpha}$** , for all  $t_0 \in (0, 1)$ ,

$$n^{(1-\alpha)/2} (\widetilde{H}_{n,\alpha}^{(IR)}(t_0) - H(t_0)) \xrightarrow[n \rightarrow \infty]{\mathcal{D}} \mathcal{N}\left(0, (\mathbb{1}_p' \Gamma^{-1}(H) \mathbb{1}_p)^{-1}\right)$$

# Definition (Stoev and Taqqu, 2006)

On pose

$$Y_{(a^+, a^-)}(t) := K(H(t)) \int_{\mathbb{R}} \frac{e^{itx} - 1}{|x|^{H(t) + \frac{1}{2}}} U_{(a^+, a^-)}(H(t), x) \widehat{W}(dx),$$

with for  $h \in (0, 1)$

$$\begin{aligned} U_{(a^+, a^-)}(h, x) &:= \frac{(a^+ e^{-i \operatorname{sign}(x)(h + \frac{1}{2}) \frac{\pi}{2}} + a^- e^{i \operatorname{sign}(x)(h + \frac{1}{2}) \frac{\pi}{2}})}{((a^+)^2 + (a^-)^2 - 2a^+ a^- \sin(\pi h))^{1/2}}. \\ &:= \frac{1}{\sqrt{\pi}} \quad \text{if } h = 1/2 \text{ and } a^+ = a^-. \end{aligned}$$

# Estimation of $H(\cdot)$

## Theorem

For a trajectory of the process  $(Y_{(a^+, a^-)}(t))_t$ ,

- If  $\max(0, 1 - 4((\eta \wedge 2) - H(t))) < \alpha < \frac{1}{2}$ ,

$$\tilde{H}_{n,\alpha}^{(IR)}(t_0) \xrightarrow[n \rightarrow \infty]{a.s.} H(t_0)$$

- If  $\max\left\{\frac{1}{1 + 2(\eta \wedge 2)}, 1 - 4(\eta \wedge 2 - H(t_0))\right\} < \alpha < 1$ ,

$$n^{(1-\alpha)/2}(\tilde{H}_{n,\alpha}^{(IR)}(t_0) - H(t_0)) \xrightarrow[n \rightarrow \infty]{\mathcal{D}} \mathcal{N}\left(0, (\mathbb{I}'_p \Gamma^{-1}(H) \mathbb{I}_p)^{-1}\right)$$

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# Generalized quadratic variations estimator of $H(\cdot)$

Define (see Istas and Lang, 1994, Benassi et al., 1998, Coeurjolly, 2005):

$$\widehat{H}_{n,\alpha}^{(QV)}(t_0) := \frac{1}{2} \frac{A'}{A'A} \left( \log \left( \frac{1}{2n^{1-\alpha}} \sum_{k=[nt_0-n^{1-\alpha}]}^{[nt_0+n^{1-\alpha}]} |X_{\frac{k+2j}{n}} - 2X_{\frac{k+j}{n}} + X_{\frac{k}{n}}|^2 \right) \right)'_{1 \leq j \leq p}$$

with  $A := (\log i - \frac{1}{p} \sum_{j=1}^p \log j)_{1 \leq i \leq p} \in \mathbb{R}^p$ .

## Theorem

*Mutatis mutandis, the limit theorems obtained for  $\widetilde{H}_{n,\alpha}^{(IR)}(t_0)$  also hold for  $\widehat{H}_{n,\alpha}^{(QV)}(t_0)$ .*

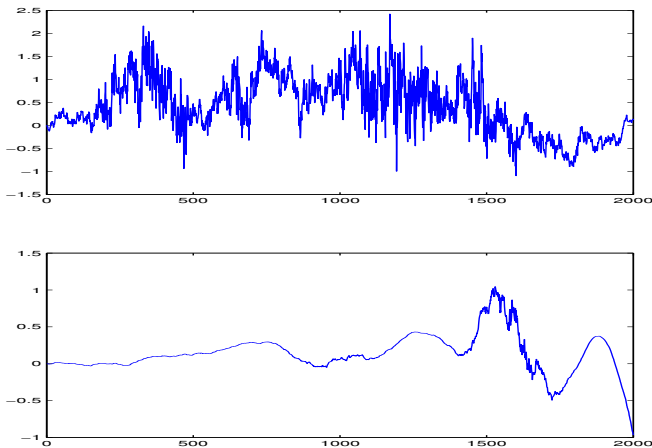
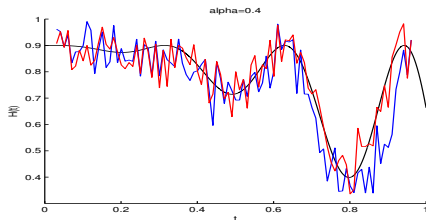
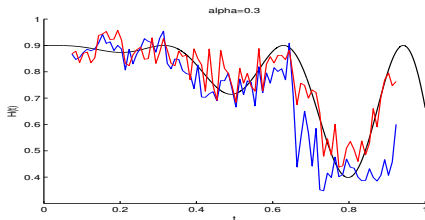
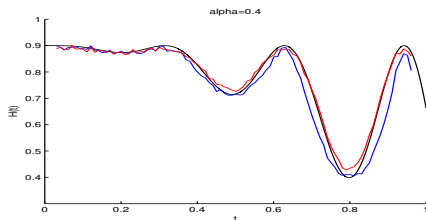
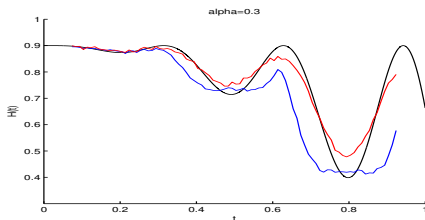


Figure: Examples of MBM trajectories (up,  $H \in \mathcal{C}^{\eta-}$  with  $\eta = 0.6$ , down  $H(t) = 0.1 + 0.8(1 - t) \sin^2(10t)$ )



**Figure:** Estimation of  $H_4(t) = 0.1 + 0.8(1-t)\sin^2(10t)$  for  $n = 6000$ ,  $\alpha = 0.3$  (left) and  $0.4$  (right)

$H_1(t) = 0.6$	$\alpha$	0.2	0.3	0.4	0.5
$n = 2000$	$\sqrt{\text{MISE}}$ for $\tilde{H}_{n,\alpha}^{(QV)}$	0.041	0.051	0.069	0.096
	$\sqrt{\text{MISE}}$ for $\tilde{H}_{n,\alpha}^{(IR)}$	0.061	0.077	0.106	0.145
$n = 6000$	$\sqrt{\text{MISE}}$ for $\tilde{H}_{n,\alpha}^{(QV)}$	0.025	0.033	0.050	0.074
	$\sqrt{\text{MISE}}$ for $\tilde{H}_{n,\alpha}^{(IR)}$	0.037	0.049	0.076	0.115
$H_2(t) = 0.1 + 0.8t$	$\alpha$	0.2	0.3	0.4	0.5
$n = 2000$	$\sqrt{\text{MISE}}$ for $\tilde{H}_{n,\alpha}^{(QV)}$	0.170	0.073	0.072	0.096
	$\sqrt{\text{MISE}}$ for $\tilde{H}_{n,\alpha}^{(IR)}$	0.059	0.071	0.098	0.135
$n = 6000$	$\sqrt{\text{MISE}}$ for $\tilde{H}_{n,\alpha}^{(QV)}$	0.114	0.044	0.048	0.070
	$\sqrt{\text{MISE}}$ for $\tilde{H}_{n,\alpha}^{(IR)}$	0.036	0.046	0.069	0.103
$H_3(t) = 0.5 + 0.4 \sin(5t)$	$\alpha$	0.2	0.3	0.4	0.5
$n = 2000$	$\sqrt{\text{MISE}}$ for $\tilde{H}_{n,\alpha}^{(QV)}$	0.362	0.123	0.080	0.096
	$\sqrt{\text{MISE}}$ for $\tilde{H}_{n,\alpha}^{(IR)}$	0.093	0.071	0.091	0.124
$n = 6000$	$\sqrt{\text{MISE}}$ for $\tilde{H}_{n,\alpha}^{(QV)}$	0.260	0.077	0.052	0.072
	$\sqrt{\text{MISE}}$ for $\tilde{H}_{n,\alpha}^{(IR)}$	0.057	0.047	0.065	0.097
$H_4(t) = 0.1 + 0.8(1 - t) \sin^2(10t)$	$\alpha$	0.2	0.3	0.4	0.5
$n = 2000$	$\sqrt{\text{MISE}}$ for $\tilde{H}_{n,\alpha}^{(QV)}$	0.320	0.164	0.117	0.112
	$\sqrt{\text{MISE}}$ for $\tilde{H}_{n,\alpha}^{(IR)}$	0.165	0.098	0.091	0.112
$n = 6000$	$\sqrt{\text{MISE}}$ for $\tilde{H}_{n,\alpha}^{(QV)}$	0.251	0.135	0.071	0.078
	$\sqrt{\text{MISE}}$ for $\tilde{H}_{n,\alpha}^{(IR)}$	0.148	0.062	0.067	0.091

Table: Estimators  $\tilde{H}_{n,\alpha}^{(QV)}$  et  $\tilde{H}_{n,\alpha}^{(IR)}$  when  $H(\cdot)$  is a  $\mathcal{C}^\infty$  function.

$H \in \mathcal{C}^{1.5}(0, 1)$	$\alpha$	0.2	0.3	0.4	0.5
$n = 2000$	$\sqrt{\widehat{MISE}}$ for $\tilde{H}_{n,\alpha}^{(QV)}$	0.261	0.112	0.085	0.098
	$\sqrt{\widehat{MISE}}$ for $\tilde{H}_{n,\alpha}^{(IR)}$	0.098	0.077	0.093	0.128
$n = 6000$	$\sqrt{\widehat{MISE}}$ for $\tilde{H}_{n,\alpha}^{(QV)}$	0.164	0.066	0.053	0.070
	$\sqrt{\widehat{MISE}}$ for $\tilde{H}_{n,\alpha}^{(IR)}$	0.054	0.047	0.066	0.098
$H \in \mathcal{C}^{0.6}(0, 1)$	$\alpha$	0.2	0.3	0.4	0.5
$n = 2000$	$\sqrt{\widehat{MISE}}$ for $\tilde{H}_{n,\alpha}^{(QV)}$	0.140	0.086	0.081	0.094
	$\sqrt{\widehat{MISE}}$ for $\tilde{H}_{n,\alpha}^{(IR)}$	0.088	0.078	0.096	0.135
$n = 6000$	$\sqrt{\widehat{MISE}}$ for $\tilde{H}_{n,\alpha}^{(QV)}$	0.130	0.067	0.056	0.071
	$\sqrt{\widehat{MISE}}$ for $\tilde{H}_{n,\alpha}^{(IR)}$	0.066	0.052	0.067	0.103

Table: Estimators  $\tilde{H}_{n,\alpha}^{(QV)}$  et  $\tilde{H}_{n,\alpha}^{(IR)}$  when  $H(\cdot)$  is a  $\mathcal{C}^{\eta-}$  function.

# Figures

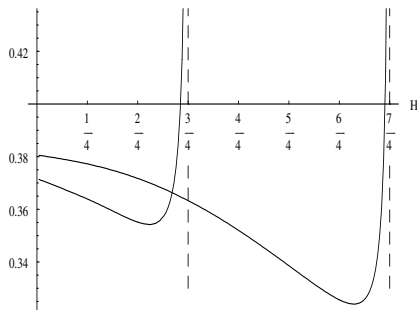
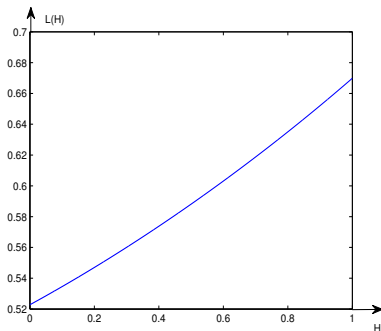


Figure: Graphs of functions  $H \mapsto \Lambda(H)$  and  $H \mapsto \sigma(H)$